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A Systematic Review and Meta-Analysis of

Workplace Mindfulness Training Randomized Controlled Trials.

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Abstract

This meta-analytic review responds to promises in the research literature and public domain about the benefits of workplace mindfulness training. It synthesizes randomized controlled trial evidence from workplace-delivered training for changes in mindfulness, stress, mental health, wellbeing and work performance outcomes. Going beyond extant reviews, this paper explores the influence of variability in workforce and intervention characteristics for reducing perceived stress. Meta-effect estimates (Hedge's *g*) were computed using data from 23 studies. Results indicate beneficial effects following training for mindfulness (g=0.45, p<0.001) and stress (g=0.56, p<0.001); for the mental health indicators anxiety (g=0.62, p<0.001) and psychological distress (g=0.69, p<0.001); and for wellbeing (g=0.46, p=0.002) and sleep (g=0.26, p=0.003). No conclusions could be drawn from pooled data for burnout due to ambivalence in results, for depression due to publication bias, or for work performance due to insufficient data. The potential for integrating the construct of mindfulness within demands-resources, coping and prevention theories of work stress is considered in relation to the results. Limitations to study designs and reporting are addressed, and recommendations to advance research in this field are made.

Keywords: mindfulness; stress; work; meta-analysis; mental health; wellbeing

Mindfulness is defined as an intentional attentiveness to present moment experience with an orientation of curiosity, openness and acceptance (Bishop et al., 2004). Over the last 30-plus years intervention research has shown mindfulness is open to development, and that established training programs result in reduced stress and improved mental health (Creswell, 2016; Gu, Strauss, Bond, & Cavanagh, 2015; Kabat-Zinn, Lipworth, & Burney, 1985; Khoury, Sharma, Rush, & Fournier, 2015; Lutz et al., 2014; Pascoe, Thompson, Jenkins, & Ski, 2017). More recent research suggests the benefits of mindfulness training may go beyond personal wellbeing to include improved work performance and relationships (Allen & Paddock, 2015; Good et al., 2016). Given the pertinence of these outcomes for working populations, interventions that develop mindfulness have gained popularity in organizations over the last decade (Lomas et al., 2017). With an estimated annual turnover of US\$1.1 billion in 2017 (Scott, 2017), and regular, largely positive media coverage (Lauricella, 2016) the mindfulness training industry is booming.

However, mindfulness training programs delivered in workplaces often vary from the training protocols upon which most scientific evidence is based (Allen et al., 2015; Hyland, Lee, & Mills, 2015). These variations include reduced time commitment (or dose) of training and the use of flexible delivery methods to meet the demands of contemporary work environments (Crane et al., 2016). The result is a heterogeneous collection of mindfulness courses, ostensibly teaching the same set of skills.

A meta-analysis gathers together results from multiple scientific studies and offers a summary of the quantitative evidence. Only one meta-analysis on this topic has been published before now (Virgili, 2015). This study looked at the single outcome of psychological distress and found overall positive effects for working adults following mindfulness training. This outcome is an important indicator of benefit, as it is a key risk

factor for mental and physical health problems that are associated with chronic stress. Virgili's work was not limited to interventions delivered within the work environment. The current paper focuses only on workplace-delivered training and supports and extends the meta-analytic evidence to include a wider range of the outcomes promised in the literature. We draw on randomized controlled trial (RCT) data, as this is currently accepted as the best way to understand if effects are attributable to treatment in intervention research (Friedman, Furberg, & DeMets, 2010). Pooling data from studies using validated outcome measures and conducted in real-world settings under everyday conditions will help clarify the Stage IV¹ evidence base for the effectiveness of mindfulness training (Dimidjian & Segal, 2015; Michalak & Heidenreich, 2018).

In this meta-analysis we ask if the best available evidence from workplace-based mindfulness training supports claims of reduced stress, and of benefits for mental health, wellbeing and work performance. Further, for the first time, we explore how workforce characteristics and variations in intervention formats influence training outcomes.

Workplace stress and related problems

Employee stress is problematic for employee mental health and wellbeing and impacts human resourcing and economic outcomes for the organizations in which they work. Occupational health psychology (OHP) research commonly draws on theoretical frameworks that attribute employee stress to the perception that work demands exceed available resources (Bakker & Demerouti, 2017). Work-related stressors include a lack of perceived control or opportunity, role conflict or ambiguity, effort-reward imbalance, isolation, uncomfortable work conditions, irregular work hours, perceived injustices and difficult relationships

¹ National Institutes of Health Stage Model for clinical research (Onken, Carroll, Shoham, Cuthbert, & Riddle, 2013)

(Hargrove, Quick, Nelson, & Quick, 2011). When such situations are perceived to be beyond coping capacity (i.e. appraised as a threat) a series of automatic cognitive and neurobiological reactions occur (Ganster, Crain, & Brossoit, 2018; Garland, Hanley, Baker, & Howard, 2017; Taren et al., 2015). If the perceived threat is not resolved, a cumulative cycle of stress reactivity can occur (Garland et al., 2017). Sustained stress depletes the body's physiological, attentional and emotional coping mechanisms, and reduces capacity to cope with future challenges. This state of chronic stress is a known contributor to clinical emotional (e.g. depression), somatic (e.g. hypertension, poor immune function) and behavioral (e.g. aggression, substance use) problems (Burton, Chen, Schultz, & Li, 2017; Cuthbert & Insel, 2013; Ganster et al., 2018; Gold, 2015).

Chronic work-related stress also negatively impacts organizations through staff disengagement, attentional deficits, absenteeism and working while unwell (presenteeism), and leads to role adjustments and workers' compensation claims (Dewa, McDaid, & Ettner, 2007; Dollard & Neser, 2013). Growing awareness of these organizational impacts, coupled with a desire to nurture employee wellbeing, are driving a rise in the uptake of work-based stress-management interventions (Memish, Martin, Dawkins, Bartlett, & Sanderson, 2017).

Firm conclusions regarding best-practice programs for workplace stress management are hard to find because relevant research is conducted in variable contexts, using different programs, outcomes and methodologies (Murphy, 1996). However, approaches that aim to improve coping skills using adaptive strategies appear to be effective (Bhui, Dinos, Stansfeld, & White, 2012). Adaptive strategies target the way demand-resource imbalances are perceived and include key skills of reperceiving (appraising stressors in different ways), and decentering (creating psychological distance from the stressor so that it can be seen within its broader context) (Folkman, 2013). Interventions with this positive orientation can build

employee coping capacity, improve mental health and wellbeing, and have been shown to support work-related performance (Bono, Glomb, Shen, Kim, & Koch, 2013; Dawkins, Martin, Scott, & Sanderson, 2015; Luthans, Avolio, Avey, & Norman, 2007; Roche, Haar, & Luthans, 2014; Sin & Lyubomirsky, 2009).

The potential of mindfulness training

Mindfulness training is a multi-modal intervention informed by the principles of positive psychology, with a central focus on skills that can reduce suffering and enable more effective coping. The evidence in support of mindfulness for beneficial outcomes is largely obtained from Mindfulness-Based Stress Reduction (MBSR) intervention studies. The MBSR training protocol originates from the principles of Eastern meditative and philosophical traditions but is presented in secular format². In the 1970's Kabat-Zinn et al. (1985) designed and tested MBSR to alleviate stress and suffering for people living with chronic pain. The program spans eight weeks and includes weekly 2.5-hour classes, a day-long retreat and prescribed homework of daily activities, including around 40 minutes of meditation. Evidence for the efficacy of MBSR for reducing stress in otherwise healthy populations is solid (Khoury et al., 2015). This well-articulated program has become the 'gold standard' for mindfulness training from which a range of derivations, or Mindfulness-Based Programs (MBPs), have emerged. Adaptations have been made to meet the needs of specific populations (e.g. employees, adolescents) and purposes (relationship enhancement, cognitive flexibility) (Chiesa & Malinowski, 2011). As with MBSR, MBPs include body scan, mindful movement and sitting meditations and are taught experientially in classes and then practiced as daily homework with the support of guided audio-tracks. MBPs are typically delivered by

² A detailed curriculum guide for MBSR is available from the University of Massachusetts Centre for Mindfulness in Medicine, Healthcare and Society at https://umassmed.edu/globalassets/center-for-mindfulness/documents/mbsr-curriculum-guide-2017.pdf

experienced mindfulness practitioners in group settings. According to guidelines for MBPs published by Crane et al. (2016), it is important to retain these elements to best facilitate the collective investigation of meditation experiences, the recognition of patterns of reactivity, and enable teacher guidance about the potential to respond differently.

Theoretical placement of mindfulness in Occupational Health Psychology (OHP)

Mindfulness training has not been extensively researched in the OHP field, and its placement within contemporary theories is emerging. Correlations with known protective resources hope, optimism and self-efficacy (Malinowski & Lim, 2015; Roche et al., 2014) indicate that being more mindful may be protective against workplace stress. This suggestion is supported by research into the mechanisms of mindfulness. For example, Garland's biobehavioral Mindfulness-to-Meaning Theory (MMT) suggests mindfulness training cultivates the adaptive coping skills of decentering and reappraisal (Garland et al., 2017). It is proposed these behavioral mechanisms interact with automatic neurobiological stress responses to intercept the cumulative cycle of reactivity associated with chronic stress. Further, the neuro-biological stress-buffering model proposed by Creswell & Lindsay (2014) suggests mindfulness meditation practice increases capacity for attentional, behavioral and emotional regulation through neuronal recruitment in brain regions responsible for these processes. In support of Creswell's model, neuro-imaging studies have shown greater mass in the pre-frontal cortex and smaller and less active amygdala in people with higher levels of mindfulness, indicating neurological evidence of improved regulatory capacity following training (Boyd, Lanius, & McKinnon, 2018; Lutz et al., 2014; Taren et al., 2015).

Aligning these mechanistic theories of mindfulness with intervention evidence and OHP theoretical models can help clarify how the training might address workplace stress. For example, a cross-sectional study of the influence of perceived autonomy and mindfulness

among nurses (Grover, Teo, Pick, & Roche, 2017) found mindfulness influenced participants' emotional regulation, which mediated improvements in perceived job control and support. The authors suggest these results show mindfulness may be a protective resource within the job demands-resources (JDR) theoretical model (Bakker & Demerouti, 2017). In addition, emerging research suggests training in mindfulness enhances the adaptive coping strategies of decentering and reappraisal (Josefsson, Lindwall, & Broberg, 2014; Keng, Choo, & Tong, 2018) indicating mindfulness training is a potentially useful stress management intervention in line with current theories of preventive stress management (Hargrove et al., 2011) and coping (Folkman, 2013). However, the extent to which mindfulness training in work settings has been tested within these three theoretical models has not previously been reported.

Mindfulness training at work

Based on the theoretical models discussed, it is proposed mindfulness training should increase participant mindfulness, and realize benefit for employee stress, mental health and wellbeing (Good et al., 2016; Hyland et al., 2015; Lomas et al., 2017). This proposition is supported by intervention meta-analyses for stress, mental health and wellbeing outcomes, including sleep (Khoury et al., 2015; Shallcross, Visvanathan, Sperber, & Duberstein, 2018; Virgili, 2015). It is further proposed mindfulness training is beneficial for work performance (Good et al., 2016; Hyland et al., 2015; Lomas et al., 2017). This is supported by studies of the association of mindfulness with work engagement (Shiba, Nishimoto, Sugimoto, & Ishikawa, 2015), burnout (N. Z. Taylor & Millear, 2016), leadership, productivity (King & Haar, 2017), empathy and perspective taking (Van Doesum, Van Lange, & Van Lange, 2013; Van Lange & Van Doesum, 2015) as well as attentional and cognitive functioning (Chiesa, Calati, & Serretti, 2011; Reb & Choi, 2014).

While the theoretical rationale of why mindfulness training might impact on work outcomes is logically defensible, the evidence appears to be based on few studies with small samples and quality limitations (Goyal et al., 2014; Kreplin, Farias, & Brazil, 2018; Rupprecht, Koole, Chaskalson, Tamdjidi, & West, 2018). Also, due to the time intensive nature of MBSR, an increasing number of MBPs promoted for workplace delivery are structured to maximize accessibility for and within organizations, with shortened classes and practice meditations and the use of flexible delivery modes (Chiesa & Malinowski, 2011; Crane et al., 2016; Van Gordon, Shonin, & Griffiths, 2015). These modifications are likely to impact the degree to which effects are realized (Crane et al., 2016; Jamieson & Tuckey, 2017). A better understanding of the relative importance of different training elements on efficacy will help guide future program modifications. Additionally, since MBSR was designed to be delivered in class-based format, the use of variable delivery modes is deserving of attention. In a recent review of online MBSR in non-work settings, effects were equivalent to face-to-face class-based training (Spijkerman, Pots, & Bohlmeijer, 2016), however flexible delivery has not been investigated for workplace-delivered MBPs. Further, the extent to which workplace context may influence training outcomes has not yet been systematically examined. For example, while health care and education samples are highly represented in mindfulness research (Allen et al., 2015; Hyland et al., 2015), the extent to which other populations benefit by comparison is worth exploring.

The Present Study

Limitations in the quality and quantity of studies, and the heterogeneity of intervention characteristics and study methods are reasons advanced by others for not synthesizing study data (Jamieson & Tuckey, 2017). However, we propose that the promise of mindfulness promulgated in the public domain should be supported by a synthesis of the

empirical evidence, and that sources of heterogeneity should be investigated for their influence on efficacy. A meta-analytic review of RCT evidence will highlight which outcomes are currently evidenced for workplace-delivered MBIs, and those which need further work, and thus guide future research and practice within the context of this fast-growth industry.

This paper therefore aims firstly to assess the effectiveness of mindfulness training delivered in the work context for employee mindfulness, stress, mental health, wellbeing and work performance; and secondly to explore the moderating role of workplace characteristics and of intervention dose, content and delivery mode.

Method

The Cochrane Handbook for Systematic Reviews of Interventions informed the methods used throughout this review (Higgins & Green, 2011). The review protocol was registered with PROSPERO in March 2016 (CRD42016036650).

Search strategy and study selection

Searches of publication databases (PubMed, Cochrane Library, Scopus, CINAHL and ProQuest), and of unpublished works (ProQuest Dissertations and Theses, WHO International Clinical Trials Register) were conducted in May 2016, and supplemented with hand searching of retrieved articles, press releases, conference abstracts, reviews and reports. The search strategy used controlled vocabulary (MeSH) and free text terms (Supplementary Table A). Two independent reviewers (LB, KM) assessed studies for inclusion, then for quality, before extracting characteristic and outcome data.

Studies published in English and fitting the following criteria were included. Intervention(s) were (1) explicitly described as mindfulness programs; (2) organized by employers and delivered for staff within the work context. Included studies (3) used an RCT

design with active or inactive comparators; and (4) reported data from any validated measure of mindfulness, stress, mental health, work performance or wellbeing. Studies were excluded if: (1) they were quasi-experimental, uncontrolled and non-randomized; (2) they taught primarily relaxation or yoga, and not mindfulness meditation; (3) interventions were therapeutic, such as dialectical behavior therapy, acceptance and commitment therapy and cognitive behavior therapy; and (4) used clinical or student populations.

Data extraction and synthesis

Characteristics of workplaces (sector, size, structure and industry), participants (occupation, education, marital status, age and sex), and interventions (mode, dose and content) were double coded by two independent reviewers, one with and one without prior content knowledge (LB, KM). The original inter-rater agreement was 70%; discrepancies were discussed and literature co-reviewed to attain 100% concordance. Original study outcomes were grouped into the five per-protocol review outcomes: mindfulness, stress, mental health, wellbeing and work performance (Supplementary Table B) and were included in meta-analyses if at least three studies reported sufficient data (Higgins & Green, 2011). Mental health measures and wellbeing measures were differentiated by the intent of the measure and the direction of results; wellbeing measures typically indicate an improvement when a higher score is returned.

Rating risk of bias

Bias in empirical research arises from a systematic error in study design, conduct or analysis, and may result in under- or over-estimation of intervention effects (Higgins & Green, 2011). In meta-analyses, rating bias risks provides an indicator of study quality, and differences can help explain heterogeneity in results. Bias risks were double coded (LB, KM) using the Cochrane Risk of Bias guidelines (Higgins et al., 2011). Individual studies were

scored for risk (1=low risk, 2=high risk, 3=unclear risk) for each of the bias categories listed above. Scores for each risk category were investigated as effect moderators. The potential for publication bias across our included studies was assessed by inspecting the distribution of points in the meta-analysis funnel plots (Higgins & Green, 2011; Higgins, Thompson, Deeks, & Altman, 2003; Kepes, Banks, McDaniel, & Whetzel, 2012). Duval and Tweedie's Trim and Fill method (Viechtbauer, 2010) was used to test the sensitivity of meta-analytic findings to publication bias, by adding potentially missing studies and recalculating the pooled effect size. Funnel plots were inspected and effect estimates adjusted following trim and fill analyses (Supplementary Figure A).

Estimates of effect

Standardized mean difference (SMD) effect estimates were calculated with random effects models and formulae provided by Fu et al. (2013). First, Cohen's *d* was computed for individual study outcomes using means, SDs and sample sizes at each time point. Inspection of outcome data revealed frequent imbalance at baseline, so *d* was computed using mean change scores by group, obtained by subtracting the baseline mean from the post-intervention mean for each group. If not reported, the SD for change scores (SD_{diff}) were imputed using the following formula (Fu et al., 2013; Higgins & Green, 2011).

$$SD_{diff} = \sqrt{SD_{BL}^{2} + SD_{PI}^{2} - 2r * SD_{BL} * SD_{PI}}$$

In this equation r is the mean correlation for within-group change from baseline to postintervention. If not provided, r was imputed from other studies reporting data from the same outcome measure. Where not available, we used the conservative value of r = 0.5 (Fu et al., 2013). After computing SMD for individual studies, Hedge's unbiased estimate (g) was used for the meta-analyses. This approach accommodates bias arising from the small number of

studies and allows pooling of data from multiple measures of the same construct (Durlak, 2009; Higgins & Green, 2011).

Interpretation of SMD effect estimates is in line with guidelines (J. Cohen, 1992), where 0.2 is a weak effect, 0.5 moderate and 0.8 strong. The direction of reported effect sizes is positive if the result indicates an improvement for the intervention. Forest plots are presented for each meta-analysis. Heterogeneity for meta-analyses is indicated by the Q statistic (Higgins et al., 2003), and I² is reported for interpretation. An I² = 75% is considered a high degree of heterogeneity, moderate when I² = 50% and low when I² = 25%. The direction and magnitude of change is reported narratively for outcomes that were not included in the meta-analyses. All analyses were conducted using R, with the *metafor* package (Viechtbauer, 2010).

Tests of moderation

The second aim of this review was to explore the influence on meta-analytic findings of bias risks, variations in the dose, content and delivery mode of MBPs, and of workforce characteristics. Moderator analyses were conducted using meta-regression when outcome data from the same measure was reported by at least 10 studies (Fu et al., 2013; Higgins et al., 2011). Hedge's *g* pooled effect estimates with 95% confidence intervals stratified by subgroups, and the amount of heterogeneity accounted for (\mathbb{R}^2) across sub groups are reported with tests of each moderator set at α =0.05.

Results

Study selection

Searches yielded 473 articles after removing duplicates, and 384 were excluded after reviewing titles and abstracts. Eighty-seven full-text papers were screened for eligibility. The PRISMA diagram (Figure 1) presents the flow of articles from initial searches to the final

inclusion of 27 papers. Two (C. Taylor et al., 2016; van Dongen et al., 2016) reported different analyses from already included studies (Roeser et al., 2013; van Berkel, Boot, Proper, Bongers, & van der Beek, 2014), meaning there were 25 separate primary studies reviewed. The authors of eight of these studies were contacted and asked to supplement published results. Ultimately, 23 studies had sufficient data for meta-analysis for at least one of the review outcomes.

Insert Figure 1: PRISMA flow diagram.

Study characteristics

Table 1 presents the characteristics of 27 included studies, 21 of which were not in the last meta-analysis (Virgili, 2015). All studies collected data pre- and post-intervention and 12 reported follow up results, with intervals ranging from six to 52 weeks. Most of the reviewed studies (k=20, 80%) compared the MBP with wait-list or treatment-as-usual comparators. Samples were mostly self-selected into the study in response to invitation campaigns (k=20), while the others were directed by their employers to participate. All included studies randomized participants to group prior to training commencement.

INSERT TABLE 1: Characteristics of the included studies

As anticipated, the mindfulness interventions were widely variable, ranging in dose from 10 minute self-guided meditations five days a week with no classes (Burnett & Pettijohn, 2015), to 42 hours' class-time over eight weeks, with 25 minutes' daily practice (Kemeny et al., 2012). While some used flexible delivery methods (e.g. online, videoconferencing, audio-tracks) (Aikens et al., 2014; Burnett & Pettijohn, 2015; Grégoire & Lachance, 2015; Prasek, 2015; Wolever et al., 2012) most were taught in face-to-face group format. Two studies provided no detail about the training protocol they investigated (McConachie, McKenzie, Morris, & Walley, 2014; Wolever et al., 2012). Meditation techniques such as body scan and breath meditation were common across the rest of the

included studies, with only two not explicitly including the body scan technique. One third of MBPs studied included a mindfulness theory component, and about half of the interventions included teachings on stress physiology. Most programs prescribed between-class meditation practice, although seven did not specify the amount (Ancona & Mendelson, 2014; Crain, Schonert-Reichl, & Roeser, 2016; Jay et al., 2015; McConachie et al., 2014; Roeser et al., 2013; Shapiro, Astin, Bishop, & Cordova, 2005; C. Taylor et al., 2016); and about half included micro-practices (brief exercises lasting between one and three minutes that can be used throughout the day to embed mindfulness into daily life). A checklist of MBP characteristics is included in Table 1.

The USA had the largest number of RCTs (n=18), followed by Canada (n=4) and one study each was published from Australia, Colombia, Denmark, Italy, Netherlands, Scotland and Taiwan. The total number of participants studied was 2,290, of which 1,086 (47%) participated in mindfulness programs and 1,204 (53%) in control conditions. Study sample sizes ranged from 18 to 257. Men were consistently underrepresented (average 15%) compared with women. The populations recruited included employees in finance and insurance (2), pharmaceutical (2), high-tech manufacturing (1) and public administration and safety (1), with the remaining studies split between education and training (12) and health and community services (9).

Risks of bias in the included studies

Concealment bias was common due to difficulties in blinding participants and teachers to treatment. Non-equivalence in baseline scores for study outcomes was observed in 12 studies, indicating the presence of selection bias. In most cases data was collected through surveys completed independently by participants, so detection bias was low. Attrition bias was present for half of the studies, with only 10 reporting results of intention to treat analyses and several omitting to report drop out or sample sizes at all time points. A high risk of

reporting bias was not observed. No significant influence on effects were observed for bias risks (Supplementary Table C).

Results of the Meta-Analysis

The meta-analytic findings for mindfulness, stress, mental health and wellbeing (Aim 1) are presented in Table 2 and Figures 3, 4 and 5. Funnel plots with trim and fill adjustments are available (Supplementary Figure A). Work performance outcomes were not sufficient for meta-analysis, so summative findings are reported narratively. Results for studies using active comparators and reporting follow-up outcomes are summarized.

INSERT TABLE 2: Meta-analysis results

INSERT FIGURES 3, 4 and 5: Meta-analysis forest plots

A consistent positive effect estimate was found for mindfulness across 12 studies using four different measures (Figure 2, Table 2). A slightly higher effect was detected for unidimensional (k=5, g=0.55, p=0.012, I²=76%) than for global scores reported from multidimensional measures (k=5, g=0.39, p<0.001, I²=0%). Three studies reported the Five Facet Mindfulness Questionnaire (FFMQ) sub-scale means, and when pooled a strong positive effect for the observe dimension was detected, while non-react improved moderately. The other facets, describe, act-aware and non-judge, returned weak, non-significant results.

A moderate reduction was found at post-intervention for perceived stress (k=13, g=0.56, p<0.001, I²=79%). This was the outcome with greatest consistency in measurement (Figure 3, Table 2). Various measures of job stress were used by four studies, and while a weak, positive trend was observed, this did not achieve significance.

Mental health outcomes with sufficient data for meta-analysis were psychological distress, depression, anxiety and burnout (Figure 4, Table 2). Several studies reported multiple mental health outcomes (e.g. depression and anxiety), so data was pooled by construct; we do not report a global mental health score. A consistent beneficial effect was

observed for psychological distress (k=8, g=0.69, p<0.001, I²=20%) and anxiety (k=4, g=0.62, p<0.001, I²=0%). Change in depressive symptoms yielded an overall positive effect (k=8, g=0.38, p=0.002, I²=48%). Results for burnout subscales were not significant, though a trend toward improvement was observed in pooled data.

Wellbeing measures include general wellbeing, health-related quality of life, sleep, fatigue/vitality, social functioning, work-life balance and satisfaction with life. Because studies often reported both sleep and wellbeing, data for sleep were pooled separately (Figure 4, Table 2). The overall mean effect across the eight studies reporting wellbeing was positive (k=8, g=0.46, p=0.002, I²=66%), and effects for sleep showed a small but consistent improvement following training (k=5, g=0.26, p=0.003, I²=0%).

Strength of Meta-Analytic Evidence

Meta-analytic results (Table 2) show heterogeneity was high for stress, depression and burnout, moderate for mindfulness and wellbeing and low for distress and anxiety. Heterogeneity can be methodological, procedural or contextual and was expected given the variability of study designs and interventions included in this review. The influence of heterogeneity on the robustness of pooled effect estimates is included in the discussion. Inspection of funnel plots revealed an uneven distribution of plot points for mindfulness, stress, depression and wellbeing, indicating the potential presence of publication bias. Trim and fill analyses (Viechtbauer, 2010) retained positive effect estimates when possibly missing studies were added to the model for all outcomes except for depression, which was reduced to non-significance (Higgins & Green, 2011) (Supplementary Figure A).

Effect Moderators

Perceived stress was assessed using the Perceived Stress Scale (PSS) by 12 studies, so was selected as the target outcome for sub-group analyses (Higgins & Green, 2011). We explored sources of heterogeneity on PSS effect estimates by examining the influence of bias

risks (Supplementary Table C) and tested the moderating effects of intervention dose, mode and content, and workforce characteristics (industry) (Table 3). Wide confidence intervals were observed for moderators that showed potential signals of influence, and none achieved significance.

INSERT TABLE 3 Moderation effects of intervention and workplace characteristics Review of Outcomes Excluded from Meta-Analysis

The results from MBP RCTs for productivity, work engagement, attention and psychosocial job quality are reported briefly and narratively because there was insufficient data for meta-analysis.

Productivity was assessed by four studies with ambivalent results. Absenteeism and presenteeism were in the positive direction at post-intervention in three studies, but results were not significant (Bartlett, Lovell, Otahal, & Sanderson, 2016; Roeser et al., 2013; Wolever et al., 2012), and no effect was observed post training or at 12 months follow up in a fourth study (van Berkel et al., 2014; van Dongen et al., 2016). Work engagement returned null results in one study (van Berkel et al., 2014; van Dongen et al., 2016), but significant positive effects in another (Aikens et al., 2014). Aikens interpreted positive changes in engagement to reflect a 20% increase in productivity and estimated a financial benefit of 20% of salary. In contrast due to negative findings for efficacy van Berkel et al. (2014) and van Dongen et al. (2016) reported a net cost for their custom MBP.

Changes in attention were measured using different methods and constructs by three studies with inconclusive results (Baccarani, Mascherpa, & Minozzo, 2013; Flook, Goldberg, Pinger, Bonus, & Davidson, 2013; Roeser et al., 2013). Psychosocial risk factors (job demand and control) were assessed quantitatively by two studies with non-significant findings (Bartlett et al., 2016; Huang, Li, Huang, & Tang, 2015) and qualitatively by a third (C. Taylor et al., 2016). The interviews in Taylor's study suggest participants reduced negative

appraisals of work stressors and increased adaptive strategies for coping with job stress. Social support was assessed by four studies using qualitative methods (Baccarani et al., 2013; Bartlett et al., 2016; Cohen-Katz et al., 2005; Moody et al., 2013), and results indicate improved work and family relationships, and that senior leadership and manager engagement may contribute to positive outcomes for participants.

Mediation analyses

Three studies conducted mediation analyses to test whether changes in outcomes could be attributed to changes in mindfulness. A strong mediation effect was found through mindfulness for perceived stress, psychological distress, mood and sleep quality (Aikens et al., 2014; Bartlett et al., 2016; Crain et al., 2016), and changes in resilience, vigor, quality of life, social functioning and job demand and control were partially mediated by changes in mindfulness (Aikens et al., 2014; Bartlett et al., 2016).

Studies with Active Control Groups

Of the 25 included studies five used active controls but no two studies used the same intervention and control design. Four studies compared effects of MBP participation with time-matched interventions and one provided information only. Compared with yoga³, effects for mindfulness were equivalent (d=0.04), but mindfulness training was superior for reducing stress (d=0.15) and depression (d=0.24) (Wolever et al., 2012). Depression, distress and anxiety improved more after mindfulness training than a leadership course (Pipe et al., 2009). Compared with a lifestyle intervention however, MBP participants reported changes in mindfulness were inconsistent, equivalent for depression and inferior for stress (Malarkey, Jarjoura, & Klatt, 2013). Compared with information about workplace stress, a positive effect was found for mindfulness, stress, distress and quality of life (Bartlett et al., 2016), but

³ Wolever et al. (2012) had two control groups; treatment-as-usual and yoga. Only the results from comparing mindfulness training with the inactive group were included in the meta-analysis.

compared with equivalent amounts of free time, effects for stress were in favour of the control group (Burnett & Pettijohn, 2015).

Studies with Follow-Up

Twelve included studies reported follow-up results and effects observed at postintervention appear to be retained at follow up. Post-intervention increases in mindfulness were retained at 11, 12 and 52 weeks (Crain et al., 2016; Grégoire & Lachance, 2015; Malarkey et al., 2013; Roeser et al., 2013) as were reductions in stress at three and 12 months (Grégoire & Lachance, 2015; Malarkey et al., 2013). Beneficial effects for psychological distress (Grégoire & Lachance, 2015; McConachie et al., 2014) depression and anxiety (Roeser et al., 2013) and wellbeing (McConachie et al., 2014) also remained stable at three months follow up. The study that reported null results for mindfulness, wellbeing, and engagement following a six-month mindfulness program saw a continuing absence of effect 12 months from baseline (van Berkel et al., 2014).

Discussion

This paper presents a systematic review and meta-analysis of data from workplacebased RCTs of mindfulness training (Aim 1). Results showed training increased mindfulness and had significant positive effects for perceived stress, psychological distress, anxiety, wellbeing and sleep, but evidence for improvements in work performance, depression and burnout was ambivalent. No significant results were observed in analyses of the influence of intervention or workplace characteristics (Aim 2).

Findings for changes in mindfulness from work-based MBPs appear weaker than obtained in studies of the gold standard MBSR program for healthy adults in non-work settings (g=0.53) (Khoury et al., 2015). This result would be weaker again if publication bias was addressed in line with trim and fill analysis, which added four studies (two positive, two

negative) into the plot, theoretically reducing the effect (g=0.29). This result indicates workplace MBP participation is effective for increasing mindfulness even with the extensive variation in training protocols.

We observed some variability in mindfulness effects by measurement instrument. The unidimensional Mindful Attention and Awareness Scale (MAAS) is recommended for use in organizational health interventions (Qu, Dasborough, & Todorova, 2015). Our pooled effect across the five studies using the MAAS was higher than the multifaceted instruments, but it also returned the largest indication of heterogeneity. The Five Facet Mindfulness Questionnaire (FFMQ) was used sufficiently frequently to look at pooled effect by dimension and these data indicate a lack of uniform improvement across mindfulness facets. This has been reported previously (Lomas et al., 2017) and has been identified for further research (Qu et al., 2015). It is feasible there is a sequential development of mindfulness qualities, but a larger sample with follow-up data is required to investigate this.

Effects for perceived stress across our included studies are in keeping with common occupational stress-management interventions such as cognitive behavioral therapy and relaxation training (Bhui et al., 2012), even when accounting for potential publication bias. Heterogeneity was high for pooled PSS results and was explored in moderation analyses (reported below). The PSS measure assesses the degree to which situations are appraised as stressful (S. Cohen, Kamarck, & Mermelstein, 1983), so these results could provide support for mindfulness training as a pathway for cultivating the adaptive strategies referred to in the theoretical biobehavioral MTM (Garland et al., 2017) and coping models (Folkman, 2013). Just three studies assessed job demand and control outcomes to examine whether mindfulness training reduces stress by improving psychosocial job quality, but these used differing approaches, reported no follow-up and results showed ambivalence. Theoretically, if mindfulness is to be considered a personal resource in the JD-R model, more evidence of a

positive relationship with other known protective characteristics such as hope, optimism and self-efficacy needs to be demonstrated, and mediation pathways tested in intervention research (Bakker & Demerouti, 2017). These outcomes were not assessed in any of the included studies, so while there is emerging theoretical and experimental support for the role of mindfulness as protective for work-related stress within the JD-R model, insufficient evidence has so far been gathered to confirm this relationship.

For mental health outcomes, our results showed the same robust post-intervention effect, with low heterogeneity and absence of publication bias for psychological distress found in the last meta-analysis on this topic (Virgili, 2015). This result is encouraging as it reinforces Virgili's findings and suggests no loss of efficacy when MBPs are conducted within the workplace context. Anxiety symptoms also responded positively to mindfulness training, with an effect estimate that is moderate to strong, significant and with low heterogeneity and publication bias. This suggests MBP participation can reduce anxiety as much as other established workplace health interventions (Martin, Sanderson, & Cocker, 2009). The collective evidence for anxiety (an automatic response to threat-perception) and psychological distress (an indicator of chronic stress) indicates mindfulness may help intercept the progressive cycle of stress and be considered a protective resource within the stress buffering, coping and biobehavioral models discussed (Creswell & Lindsay, 2014; Folkman, 2013; Garland et al., 2017). However, outcomes that can verify the mediating role of mindfulness for increasing adaptive coping strategies such as reperceiving and decentering were not assessed in any of the included RCTs.

While a reduction in depression symptoms was observed in the meta-analysis, heterogeneity was present, and once the potential influence of publication bias was accounted for, the estimate of effect reduced to non-significance. It is feasible the general level of depressive symptoms was low even if anxiety and distress were elevated, and this may

explain why the effects were weaker for depression among our sample than from studies in clinical settings, where depressive symptoms have consistently shown improvement (Flaxman & Bond, 2010). Correlational studies indicate burnout symptoms reduce when mindfulness increases (N. Z. Taylor & Millear, 2016). The burnout instrument used in all our included studies was the Maslach Burnout Inventory and the inconsistency in results we observed appears common for this measure in intervention research (Morse, Salyers, Rollins, Monroe-DeVita, & Pfahler, 2012). Burnout is increasingly recognized as a clinical condition requiring individualized treatment (Kakiashvili, Leszek, & Rutkowski, 2013) and, like depression, may respond better to clinically directed therapy than to a workplace training program.

Wellbeing indicators including life-satisfaction, work-life balance, fatigue and vitality showed promise in individual studies, as did data from quality of life measures. We pooled data for these outcomes and found a positive effect estimate in keeping with MBP studies in non-work settings (De Vibe, Bjørndal, Tipton, Hammerstrøm, & Kowalski, 2012; Gu et al., 2015), and a small likelihood of publication bias. Wellbeing is presented as a flow-on outcome in the model by Good et al. (2016) and its improvement across studies is encouraging. However, more RCT evidence is needed to defend claims of efficacy for each of these wellbeing outcomes independently. Our results for sleep showed improvement in keeping with evidence of mindfulness training in other contexts (Shallcross et al., 2018). As sleep quality is a known contributor to general wellbeing and occupational functioning (Burton et al., 2017; Knudsen, Ducharme, & Roman, 2007) these results support the potential for mindfulness for stress coping and recovery.

Despite relevance to the workplaces in which the studies were conducted, performance outcomes such as productivity, engagement and attention were inconsistently and infrequently assessed, and results were ambivalent. Productivity may have achieved

significance if we pooled results but this was not done due to construct and methodological variability. Presenteeism and absenteeism are important contributors to organizational productivity, but successful workplace preventive stress management interventions should also realize change on economic indicators of health-care utilization and direct costs to employers such as role accommodations and compensation claims (Hargrove et al., 2011). Change in these outcomes can take time to manifest and is likely to be more evident with longer term follow-up than immediately post-intervention. Surprisingly, work engagement was infrequently studied in the research reviewed despite the outcome being pertinent to staff performance and turnover, and informative for organizational perspectives of MBI efficacy.

The available evidence of change in attentional capacities is insufficient to compare with positive findings from MBP participation in non-work settings (Chiesa et al., 2011). This may be because testing requires a degree of manipulation and laboratory assessments that are challenging for workplace-based research. As attentional control, stability and efficiency are considered key mechanisms of mindfulness training (Good et al., 2016; Hölzel et al., 2011), and the potential benefits of attention for work performance are evident, suitable assessment techniques for use in workplace research are needed to provide evidence in support of the promise presented in the literature.

We found no RCT evidence supporting mindfulness training for leadership or creativity, decision making, citizenship behaviors, deviance or safety despite the promising articles about mindfulness for work cited in the introduction (Good et al., 2016; Hyland et al., 2015). Our findings illustrate significant gaps recently identified in the evidence for organizational outcomes (Rupprecht et al., 2018) and support critiques of the conduct and reporting of mindfulness research more generally (Jamieson & Tuckey, 2017; Lomas et al., 2017).

Detrimental effects

We expected lower effects when active controls were used (Meinert, 2012), and this may partly explain increases in stress for mindfulness participants compared with controls in two studies (Burnett & Pettijohn, 2015; Malarkey et al., 2013). However, two other studies also found detrimental effects for emotional exhaustion (Moody et al., 2013) and wellbeing (van Berkel et al., 2014) when compared with inactive controls. Two of these four studies showing some decrement following training expressly targeted employees who were identified as 'at risk' (Burnett & Pettijohn, 2015; Malarkey et al., 2013), but so did two studies that returned positive results (Huang et al., 2015; Mackenzie, Poulin, & Seidman-Carlson, 2006). The interventions assessed by Malarkey et al. (2013) and van Berkel et al. (2014) required a commitment of more than 10 hours class-time over an eight-week period, plus 20-30 minutes' daily homework. Van Berkel's custom intervention was integrated with a broader year-long health promotion campaign, and Malarkey's assessment protocol was complex, with the collection of biomarkers and survey data at multiple time points over a sixmonth period. It appears work release was not provided for participants in any of the reviewed studies and it is plausible that for these more intensive interventions, the overall demands of training when added to existing workloads served to increase rather than decrease employee stress. In contrast, Burnett's (2015) intervention was comparatively low in dose, with daily 10-minute meditation practices done at the participants' computer during work time and no classes. The study authors suggest the increase in stress in Burnett's participants may be linked to pressure to improve performance. It may also be due to a perception of additional demands amongst course participants, when compared with a control group that was given free time. These exceptions highlight the potential for benefits to be offset by the demands of this type of training.

The influence of workplace and intervention characteristics

None of the tests of moderation returned significant results (Table 3). For a handful of explanatory variables, this non-significance was more likely due to inadequate power than to the absence of influence. Acknowledging the lack of significance, we present the following brief discussion for future hypothesis generation.

Intervention Characteristics

The equivalent efficacy of flexible delivery and face-to-face classes has been observed previously in studies of online MBSR in non-work settings (Spijkerman et al., 2016). The additional control over access to training offered by flexible delivery may balance out the benefits of supportive group dynamics (Crane et al., 2016). It is also possible the nonspecific effects of group membership may be more complex in workplace interventions, where learning is with co-workers, compared with community settings where pre-existing relationships between participants are less likely (Davidson & Kaszniak, 2015). Our findings for the influence of dose are similar to those observed by Carmody and Baer (2009), and it is not yet possible to derive a minimum required dose for positive effect from the current evidence base. Including micro-practices and mindfulness theory in training protocols appears to yield some influence on reductions in stress, while yoga and stress physiology elements showed no sign of moderation. The inclusion and relative importance of specific elements of training has not previously been reported for workplace MBPs.

Workforce Characteristics

Equivalent results were observed for people across industries, suggesting this approach may be suitable for workforces beyond the caring professions. However population samples were drawn from large organizations with predominantly white-collar workforces, and so it is not clear whether mindfulness training is suitable and efficacious for other settings and role types.

Practical implications

Workplace-delivered mindfulness training programs can help employees reduce stress and improve their mental health and wellbeing. The effects of training can endure for at least 12 months. Low dose interventions that use either flexible delivery or a class-based approach appear effective, and benefit does not appear to be limited to education and health care professionals. At present it is not known whether the effectiveness of training differs by role type (e.g. blue collar, administrative, professional). The promise in the public domain of mindfulness for organizational performance outcomes is not yet supported by a quality evidence base.

While the results of this meta-analysis are positive for stress and related mental health and wellbeing, the findings regarding effects beyond personal mental health provide an important counter-point to highly cited articles (e.g. Good et al., 2016) and governmentendorsed recommendations such as those in the Mindful Nation UK report (Loughton & Morden, 2015) and the ensuing Building the Case for Mindfulness in the Workplace (Bristow, 2016). Given the considerable organizational expenditure and media coverage of the promise of mindfulness training for improving work performance, we feel this evidence synthesis, and its conclusion that these claims are ahead of the evidence, is both warranted and timely.

Limitations and directions for future research

Previous reviews have identified methodological limitations and multiple sources of heterogeneity in workplace stress management and mindfulness research literature. Our findings support these critiques and provide statistical evidence of their influence on results. Our tests for publication bias suggest there may be a 'file drawer' problem for studies that do not report significant findings, particularly for those measuring depression.

Moderation analyses could include only the 12 RCTs reporting PSS data, and so our sub-group analyses were limited. We were unable to investigate if people in certain roles

benefit more than others, due to the absence of this information in the reviewed manuscripts. Given the predominance of studies in large organizations and white collar workforces we cannot generalize to employee samples from small and medium size businesses or labor-intensive industries. The articulation of MBP format varied across studies, from little to no detail to sufficient for replication. Studies that cited external references for program format make comparing intervention characteristics less accessible for readers (McConachie et al., 2014). One study included in moderator analyses did not provide program content due to proprietary restrictions (Wolever et al., 2012).

In addition to limitations in reporting for our effect moderators, several studies did not provide detail about attrition, or report sample sizes for each time point, and while intentionto-treat analyses can account for missingness, this was not always reported either. Adherence to training protocols was also reported variously, with a small number of studies considering the influence of dose received versus dose administered, but most drew on subjective reflections and the range in measurement approaches for adherence meant these data could not be used for reporting or analysis. Follow-up data was only reported by a handful of the included studies, and frequently the time points were set for the end of a wait-list control's intervention period. This is pragmatic for conducting research in workplace settings, but better evidence would be generated by retaining control conditions through all time points and extending the follow up to one or two years from baseline. This is particularly true for organizational outcomes like productivity that may take more time to manifest than changes in personal mental health. The use of established stress management interventions as active controls in future studies would be informative for assessing the comparative benefits of mindfulness training and support organizational investment choices. Follow-up data would also generate defensible evidence of the lasting effects of training and explore the sequential development of different aspects of mindfulness. Follow-up data is also vitally important for

economic analyses, another key consideration for organizational decisions regarding training investment. The current state of evidence is insufficient for this purpose.

Process and contextual factors related to being in a workplace setting may also have influenced results (Nielsen & Miraglia, 2016), but only two studies explicitly discussed these considerations. Procedural factors such as the influence of teacher experience, self-selection versus targeted recruitment, the nature and extent of assessments and flexibility of class times and class sizes would add to the quality and depth of evidence going forward. Guidelines presented by Crane and Hecht (2018) should be followed in future to enable the thorough examination of intervention components and their relationship to results.

Despite the theoretical promise, more work is needed to explore the potential of mindfulness as a personal resource for ameliorating the demands of work (Bakker & Demerouti, 2017). Studies that examine if training in mindfulness directly influences adaptive coping strategies and/or cultivates a protective personal resource that is additional to optimism and self-efficacy will make valuable contributions to the field. Data from multiple studies using appropriately sensitive and validated measures of occupational health and performance indicators including engagement, citizenship, sociality, creativity, leadership and safety are needed to qualify claims in the public domain regarding the effects of mindfulness training for work performance.

Finally, while our results indicate that mindfulness training can be beneficial, the negative results of a small number of studies suggest potential benefits may be offset by the addition of training demands to workloads (Burnett & Pettijohn, 2015; Malarkey et al., 2013; Moody et al., 2013) and a deeper understanding of this relationship, combined with evidence about the most instrumental elements of MBPs would guide future implementation strategies.

Conclusions

Workplace-delivered mindfulness training can cultivate employee mindfulness, reduce perceived stress, anxiety and psychological distress, and be beneficial for wellbeing and sleep quality. Effect sizes are in keeping with other well recognized stress-management interventions like relaxation and CBT.

The promise for enhanced work performance following mindfulness training is not yet supported by the evidence, and claims of improvements in organizational citizenship, leadership, deviance, safety or creativity cannot be defended at the present time with RCT evidence.

We recommend future studies use validated measures of performance for organizational and individual outcomes, conduct follow-up assessments, replicate interventions in different settings, and continue to explore effects for work sectors beyond education and health. Future researchers are encouraged to account for baseline imbalance, use comparators that can inform investment decisions, include economic evaluations, conduct mediation analyses and report potential moderators such as those used in our exploratory analyses. Empirical studies that examine workplace-based mindfulness training within theoretical stress buffering, biobehavioral and JD-R frameworks are needed to place mindfulness training defensibly within the occupational health psychology literature.

While we can conclude from the current study that workplace-delivered MBPs can effectively reduce employee stress and related problems, addressing the identified limitations will help clarify the relative efficacy of differing training approaches for individuals, and ensure the promised benefits of mindfulness training for organizational outcomes are evidence based.

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Tables and Figures for inclusion in the published paper.

Table 1: Review of Workplace Mindfulness RCTs: Characteristics of Included Studies

Table 2: Review of Workplace Mindfulness RCTs: Meta-Analytic Effect Estimates

 Table 3: Review of Workplace Mindfulness RCTs: Moderation Effects of Intervention and Workplace Characteristics on Perceived Stress

Figure 1. Workplace Mindfulness RCTs: PRISMA Flow Diagram

Figure 2. Workplace Mindfulness RCTs: meta-analysis results for mindfulness

Figure 3. Workplace Mindfulness RCTs: meta-analysis results for perceived stress and job stress

Figure 4. Workplace Mindfulness RCTs: meta-analysis results for mental health and wellbeing

Supplementary materials

Table A. Systematic Review of Workplace Mindfulness RCTs: Search Strategy

Table B. Systematic Review of Workplace Mindfulness RCTs: Study Outcomes Grouped into Review Outcomes

Table C. Meta-Analysis of Workplace Mindfulness RCTs: Moderation Effects of Bias Risks on Perceived Stress

Figure A. Workplace Mindfulness RCTs: meta-analysis funnel plots with SMD effect estimates with trim and fill adjustments for publication bias

| study | Intervention | Comparison | Delivery mode | S | class# | Baily homework# | Mindful theory | Stress physiology | scan | th | Open monitoring | ing | | Journaling | Micro-practice | Compassion | <u>5</u> 0 | |
|----------------------------|---|-------------|------------------------------|---------|----------------------------|-----------------|----------------|-------------------|-------------|----------|-----------------|-----------|--------|------------|-----------------------|------------|------------|--|
| Study Z Aikens, 2014 89 | Custom online mindfulness | MC Com | Delix Online supported | L Weeks | 99 Each class [#] | 05 Daily | Mind | Stres | × Body scan | × Breath | x Open | × Walking | × Yoga | Jouri | x Micro | Com | Eating | Population Pharmaceutical manufacturing |
| Ancona, 2014 43 | | TAU | F2F class | 3 | 45 | ns | х | х | | х | | | х | | | | | School teachers |
| Baccarani, 2013 20 | | TAU | F2F class | 4 | 90 | 45 | | | | х | х | | х | | х | | | University staff |
| Bartlett, 2015 12 | | Information | F2F class | 5 | 90 | 20 | X | X | X | X | | | | | X | | | Public servants |
| Burnett, 2015 55 | 5 MBST | Free-time | Self- guided audio | 5 | 10 | 10 | | | x | x | | | | | | | | Health service employees |
| **Cohen-Katz, 27 2015 | 7 MBSR | WL | F2F class | 8 | 150 | 45 | X | X | X | X | X | X | X | X | X | | | Nurses |
| Crain, 2016 11 | 13 Workplace Mindfulness Training | WL | F2F class | 8 | v | ns | X | X | X | х | X | | | | x | x | | School teachers |
| **Duchemin, 32 2015 | | WL | F2F class | 8 | 60 | 20 | | | x | х | | | x | | | | х | Hospital-based intensive care |
| Flook, 2013 18 | 8 MBSR | WL | F2F class | 8 | 150 | 30 | х | х | х | х | х | х | х | | х | х | Х | School teachers |
| Gregoire, 2015 49 | 9 Custom MBP | WL | Seminar + self-guided | 5 | 60 | 15 | | | X | X | X | | | | | x | | Call center staff |
| Huang, 2015 14 | 44 MBSR-ld | TAU | F2F class | 8 | 120 | 45 | х | x | х | X | X | X | X | | х | x | | Manufacturing admin & production staff |
| Jay, 2015 11 | 12 Custom MBCT-ld | TAU | F2F class | 10 | 50 | ns | | | х | х | | | х | | | | | Pharmaceutical lab technicians |

Table 1: Review of Workplace Mindfulness RCTs: Characteristics of Included Studies

| Kemeny, 2012 | 82 | Custom Contemplati ve Training | WL | F2F class | 8 | 480 | 25 | | X | X | X | X | Х | х | | | х | | School teachers |
|----------------------|-----|--------------------------------------|---------------------|-----------------------|----|-----|----|---|---|---|---|---|---|---|---|---|---|---|----------------------------------|
| Klatt, 2009 | 48 | MBSR-ld | WL | F2F class | 8 | 60 | 20 | х | Х | х | Х | Х | х | х | Х | | | | University staff |
| Mackenzie, 2006 | 30 | MBSR-ld | WL | F2F class | 4 | 30 | 10 | х | | х | х | X | | х | | X | | х | Nurses |
| Malarkey, 2013 | 186 | MBP-ld | Lifestyle education | F2F class | 8 | 60 | 20 | | х | х | х | | X | х | х | х | | | University staff |
| Manotas, 2014 | 83 | MBSR-ld | TAU | F2F class | 4 | 120 | 25 | X | X | X | Х | Х | | X | | | | х | Hospital-based health care staff |
| McConachie, 2014 | 120 | ACT plus Mindfulness | WL | F2F class | 6 | 480 | ns | | | | | | | | х | | | | Intellectual disability carers |
| Moody, 2013 | 48 | MBSR- custom | TAU | F2F class | 8 | 60 | 20 | Х | х | Х | Х | х | | х | Х | х | Х | Х | Oncology staff |
| Pipe, 2009 | 33 | MBSR-ld | Leadership course | F2F class | 4 | 120 | 30 | | X | X | Х | | | X | | Х | | | Nurse leaders |
| Prasek, 2015 | 192 | Sherman Project | WL | Online self-guided | 7 | 15 | 10 | | х | х | х | | | х | х | х | | х | University staff |
| Roeser, 2013 | 113 | SMART-in- Education | WL | F2F class | 8 | 150 | ns | | X | X | х | X | x | X | X | X | x | х | School teachers |
| Shapiro, 2005 | 38 | MBSR | WL | F2F class | 8 | 120 | ns | | | х | Х | X | | X | | X | X | | Hospital-based health carers |
| *Taylor, 2016 | | SMART-in- Education | WL | F2F class | 9 | 150 | ns | | | | | | | | | | | | School teachers |
| van Berkel, 2014 | 257 | Mindful VIP (custom) | TAU | F2F class | 8 | 90 | 30 | | | Х | Х | | Х | X | X | X | | Х | Research institute |
| *van Dongen, 2016 | | Mindful VIP (custom) | TAU | F2F class | 8 | 90 | 30 | | | | | | | | | | | | Research institute |
| Wolever, 2012 | 239 | Mindfulness at Work | Yoga & TAU | Online supported | 12 | 60 | 10 | | | | | | | | | | | | Insurance |

* Reports additional findings from primary study; ** Not included in meta-analysis; # Units reported in hours; ns = not specified; x = included in description of program WL = wait-list control group; TAU = treatment as usual control group; F2F = face to face

MBSR (ld)= Mindfulness-Based Stress Reduction (low-dose); MBST = Mindfulness-Based Stress Training; ACT = Acceptance and Commitment Training

| | | SMD | 95% | 6 CI | Significance | Heterog | geneity |
|------------------------|----|-----------|-------|-------|--------------|---------|----------------|
| | k | Hedge's g | Lower | Upper | р | Q | \mathbf{I}^2 |
| MINDFULNESS | | | | | | | |
| All data | 10 | 0.45 | 0.26 | 0.64 | < 0.001 | 19.52 | 54% |
| MAAS | 5 | 0.55 | 0.12 | 0.98 | 0.012 | 16.54 | 76% |
| Multi-facet measures | 5 | 0.39 | 0.23 | 0.55 | < 0.001 | 2.92 | 0% |
| FFMQ Observe | 3 | 0.82 | 0.37 | 1.26 | < 0.001 | 3.96 | 49% |
| FFMQ Describe | 3 | 0.23 | -0.19 | 0.64 | 0.280 | 3.45 | 42% |
| FFMQ Act Aware | 3 | 0.07 | -0.23 | 0.36 | 0.646 | 0.78 | 0% |
| FFMQ Non-react | 3 | 0.52 | 0.23 | 0.82 | 0.001 | 1.85 | 0% |
| FFMA Non-judge | 3 | 0.28 | -0.01 | 0.58 | 0.059 | 0.71 | 0% |
| STRESS | | | | | | | |
| All data | 13 | 0.56 | 0.29 | 0.83 | < 0.001 | 56.15 | 79% |
| PSS | 12 | 0.54 | 0.26 | 0.83 | 0.00 | 54.77 | 80% |
| Job-stress | 4 | 0.10 | -0.11 | 0.32 | 0.339 | 3.88 | 23% |
| MENTAL HEALTH | | | | | | | |
| Psychological distress | 8 | 0.69 | 0.49 | 0.90 | < 0.001 | 8.73 | 20% |
| Depression | 8 | 0.38 | 0.14 | 0.62 | 0.002 | 13.53 | 48% |
| Anxiety | 4 | 0.62 | 0.32 | 0.92 | < 0.001 | 0.77 | 0% |
| Burnout - EE | 4 | 0.52 | 0.19 | 1.22 | 0.149 | 11.41 | 74% |
| Burnout - PA | 3 | 0.46 | -0.03 | 0.96 | 0.066 | 2.69 | 26% |
| Burnout - DP | 3 | 0.16 | 0.25 | 0.58 | 0.439 | 0.25 | 0% |
| WELLBEING | | | | | | | |
| All data | 8 | 0.46 | 0.17 | 0.72 | 0.002 | 20.39 | 66% |
| General wellbeing | 3 | 0.32 | 0.01 | 0.63 | 0.046 | 3.04 | 34% |
| Sleep Problems | 5 | 0.26 | 0.09 | 0.43 | 0.003 | 0.96 | 0% |

SMD: Standardized Mean Difference effect estimates computed using a random effects model with α =0.05; MAAS: Mindful Attention and Awareness Scale; FFMQ: Five Facet Mindfulness Questionnaire; PSS: Perceived Stress Scale (10 or 14-item versions); Burnout - EE: Emotional Exhaustion; Burnout - PA: Personal Accomplishment; Burnout - DP: Depersonalization.

| Intervention and Work | | | | | ed Stress | | |
|------------------------------|----------------|-------------|----------------|-------------------------|-----------|-------|------|
| | No. studies | Tes mode | t of rators | Sub- group Hedges | 959 | %CI | |
| | k | р | \mathbb{R}^2 | g | Lower | Upper | Z |
| Perceived Stress | 12 | - | | 0.54 | 0.26 | 0.83 | |
| Influence of Intervention Ch | aracterist | ics | | | | | |
| Delivery mode | | 0.557 | 0% | | | | |
| F2F weekly classes | 8 | | | 0.54 | 0.18 | 0.91 | 2.91 |
| Other format | 4 | | | 0.56 | 0.01 | 1.01 | 2.40 |
| Dose (contact time) | | | | | | | |
| Class hours | | 0.520 | 0% | | | | |
| Up to 7 | 4 | | | 0.58 | 0.08 | 1.08 | 2.26 |
| 8 and over | 8 | | | 0.53 | 0.17 | 0.89 | 2.86 |
| Weeks | | 0.489 | 0% | | | | |
| Up to 7 | 5 | | | 0.66 | 0.24 | 1.08 | 3.09 |
| 8 and over | 7 | | | 0.46 | 0.08 | 0.84 | 2.38 |
| Dose (homework) | | | | | | | |
| Daily home practice | | 0.806 | 0% | | | | |
| Up to 10 minutes | 5 | | | 0.47 | 0.07 | 0.88 | 2.28 |
| More than 10 minutes | 7 | | | 0.60 | 0.19 | 1.02 | 2.84 |
| Content* | | | | | | | |
| Stress physiology | | 0.783 | 0% | | | | |
| Included | 7 | | | 0.53 | 0.15 | 0.90 | 2.73 |
| Not included | 4 | | | 0.44 | 0.01 | 0.89 | 1.90 |
| Mindfulness theory | | 0.079 | 27% | | | | |
| Included | 5 | | | 0.73 | 0.40 | 1.07 | 4.37 |
| Not included | 6 | | | 0.29 | 0.07 | 0.65 | 1.56 |
| Movement/yoga | | 0.806 | 0% | | | | |
| Included | 9 | | | 0.47 | 0.19 | 0.75 | 3.28 |
| Not included | 2 | | | 0.53 | -0.82 | 1.89 | 0.77 |
| Micro-practice | | 0.142 | 19% | | | | |
| Included | 6 | | | 0.65 | 0.37 | 0.92 | 4.64 |
| Not included | 5 | | | 0.28 | -0.18 | 0.75 | 1.19 |
| Influence of Work Character | ristics | | | | | | |
| Industry | | 0.296 | 30% | | | | |
| Human services | 8 | | | 0.50 | 0.10 | 0.90 | 2.47 |
| Other | 4 | | | 0.62 | 0.26 | 0.99 | 3.34 |

Table 3. Review of Workplace Mindfulness RCTs: Moderation Effects of Intervention and Workplace Characteristics on Perceived Stress

Hedge's g: standardized mean difference by group for perceived stress stratified by moderator sub group;

F2F: face-to-face; R²: amount of heterogeneity accounted for; p: test of moderators (α =0.05)

* K=11 for sub-group analyses by content, as one of the studies did not report this information.

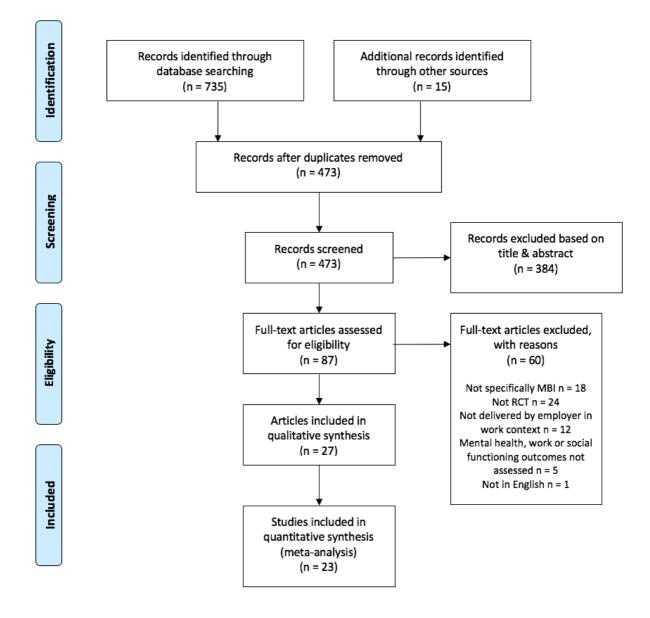


Figure 1. Workplace Mindfulness RCTs: PRISMA Flow Diagram

| Author, Year | Instrument | | Effect size [95% CI] | Weight |
|---------------------------|------------------|---------------------|------------------------|------------------------|
| Mindfulness (glob | bal scores) | | | |
| Bartlett, 2015 | MAAS | ; ⊢_ ∎{ | 0.97 [0.45 – 1.49] | 7.9% |
| Crain, 2016 | FFMQ | ; ⊢∎ (| 0.55 [0.16 – 0.94] | 10.6% |
| Gregoire, 2015 | MAAS | | 1.06 [0.41 – 1.71] | 5.9% |
| Kemeny, 2012 | MAAS | ├── ■──┤ | 0.43 [-0.01 - 0.87] | 9.5% |
| Klatt, 2009 | MAAS | F <u>+</u> −−−−1 | 0.50 [-0.12 - 1.12] | 6.3% |
| Malarkey, 2013 | TMS | l÷-■1 | 0.21 [-0.07 - 0.49] | 13.8% |
| Manotas, 2014 | FFMQ | ; ⊢ 1 | 0.48 [0.04 – 0.92] | 9.5% |
| Roeser, 2013 | FFMQ | | 0.53 [0.14 – 0.92] | 10.6% |
| Van Berkel, 2014 | MAAS | ⊢-∰1 | 0.00 [-0.28 - 0.28] | 13.8% |
| Wolever, 2012 | CAMS-R | | 0.38 [0.04 – 0.72] | 12.0% |
| | Pooled effect | • | 0.45 [0.26 – 0.64] | |
| Mindfulness (FFN | IQ facet scores) | | | |
| Observe | | | | |
| Aikens, 2014 | | ; ⊢ ∎1 | 1.17 [0.73 – 1.61] | 41.5% |
| Flook, 2013 | | ⊢ | 0.31 [–0.63 – 1.25] | 17.1% |
| Manotas, 2014 | | i ⊢∎1 | 0.67 [0.23 – 1.11] | 41.5% |
| | Pooled effect | | 0.82 [0.37 – 1.26] | |
| Describe | | - | | |
| Aikens, 2014 | | ∳ ∎I | 0.43 [–0.01 – 0.87] | 42.2% |
| Flook, 2013 | | <u>⊢ ;</u> | 0.57 [-0.37 - 1.51] | 15.6% |
| Manotas, 2014 | | ⊢ | -0.10 [-0.54 - 0.34] | 42.2% |
| <u>.</u> . | Pooled effect | | 0.23 [-0.19 - 0.52] | |
| Act-aware | | | | |
| Aikens, 2014 | | | 0.21 [-0.23 - 0.65] | 45.1% |
| Flook, 2013 | | | 0.06 [-0.88 - 1.00] | 9.8% |
| Manotas, 2014 | | } ≣ ¦1 | –0.07 [–0.51 – 0.37] | 45.1% |
| Non read | Pooled effect | \bullet | 0.07 [-0.23 - 0.36] | |
| Non–react Aikens, 2014 | | i | 0.75 [0.31 – 1.19] | 45 10/ |
| Flook, 2013 | | | 0.34 [-0.60 - 1.28] | 45.1% 9.8% |
| Manotas, 2013 | | | 0.34 [-0.10 - 0.78] | 9.8% 45.1% |
| Manolas, 2014 | | | | 45.1% |
| Non-judge | Pooled effect | | 0.52 [0.23 – 0.82] | |
| Aikens, 2014 | | | 0.19 [-0.25 - 0.63] | 45.1% |
| Flook, 2013 | | | 0.09 [-0.85 - 1.03] | 4 <u>3.</u> 1% 9.8% |
| Manotas, 2014 | | · · · · · | 0.42 [-0.02 - 0.86] | 45.1% |
| | Pooled effect | | 0.28 [-0.01 - 0.58] | 45.178 |
| | Fooled effect | | 0.20 [-0.01 - 0.06] | |
| | | | | |
| | | -1 -0.5 0 0.5 1 1.5 | 2 | |
| | | | - | |

Figure 2. Workplace Mindfulness RCTs: meta-analysis results for mindfulness

Standardized Mean Difference

Fig 2 Legend: MAAS: Mindful Attention and Awareness Scale; FFMQ: Five Facet Mindfulness Questionnaire; TMS: Toronto Mindfulness Scale; CAMS-R: cognitive and Affective Mindfulness Scale Revised

| Author, Year | Instrument | Effect size [95% CI] | Weight |
|------------------|---------------|------------------------------------|--------|
| Perceived Stress | S | | |
| Aikens, 2014 | PSS-14 | ▶ ■ 0.84 [0.32 – 1.36] | 8.1% |
| Bartlett, 2015 | PSS-14 | ▶ 1.21 [0.27 – 2.15] | 7.5% |
| Burnett, 2015 | PSS-10 | -0.17 [-0.79 - 0.45] | 6.5% |
| Gregoire, 2015 | PSM-9 | →−−− → 0.84 [0.50 – 1.18] | 6.7% |
| Huang, 2015 | PSS-10 | 0.38 [-0.10 - 0.86] | 8.9% |
| Jay, 2015 | PSS-14 | 0.26 [-0.08 - 0.60] | 8.5% |
| Klatt, 2009 | PSS-10 | 0.60 [-0.13 - 1.33] | 6.7% |
| Malarkey, 2013 | PSS-10 | -0.21 [-0.99 - 0.57] | 9.3% |
| Manotas, 2014 | PSS-10 | ▶ ■ 0.99 [0.09 – 1.89] | 7.8% |
| Moody, 2013 | PSS-14 | 0.56 [0.08 – 1.04] | 6.9% |
| Prasek, 2015 | PSS-10 | 0.36 [0.08 – 0.64] | 8.9% |
| Shapiro, 2005 | PSS-10 | • 0.83 [0.39 – 1.27] | 5.4% |
| Wolever, 2012 | PSS-10 | ▶ ■ 1.02 [0.40 – 1.64] | 8.9% |
| | Pooled effect | • 0.56 [0.29 – 0.83] | |
| Job Stress | | | |
| Ancona, 2014 | TSI | 0.54 [-0.17 - 1.25] | 10.5% |
| McConachie, 2014 | SSQ | 0.02 [-0.53 - 0.57] | 28.6% |
| Roeser, 2013 | TSI | 0.26 [-0.08 - 0.60] | 23.0% |
| Van Berkel, 2014 | NFR | -0.05 [-0.95 - 0.85] | 37.9% |
| | Pooled effect | 0.10 [-0.11 - 0.32] | |
| | | | |
| | | -1 -0.5 0 0.5 1 1.5 2 | |
| | | Standardized Mean Difference | |

Figure 3. Workplace Mindfulness RCTs: meta-analysis results for perceived stress and job stress

Fig 3 Legend: PSS: Perceived Stress Scale (14 or 10-item); PSM-9: Psychological Stress Measure (9 item); TSI: Teacher Stress Inventory; SSQ: Staff Stressor Questionnaire; NFR: Need for Recovery subscale from Questionnaire on the Experience and Evaluation of Work

| Figure 4. Workplace Mindfulness | RCTs: meta-analysis results for mental health and |
|---------------------------------|--|
| wellbeing | |

| Author, Year | Instrument | Effect size [95% C | l] Weigh |
|-----------------------------------|---------------|---|----------|
| Psychological Distre | | | |
| | | | 1 10.10 |
| Bartlett, 2016 | K-10 | | |
| Flook, 2013 | SCL-90-R-GSI | | |
| Gregoire, 2015 | PDMS | | - |
| Huang, 2015 | CHQ-12 | 0.87 [0.53 – 1.21 | - |
| Manotas, 2014 | BSI–GSI | | |
| McConachie, 2014 | GHQ-12 | .35 [0.01 – 0.69 | |
| Pipe, 2009 | SCL-90-R-GSI | ; | |
| Shapiro, 2005 | BSI–GSI | 0.47 [-0.31 - 1.25 |] 5.6% |
| Depression | Pooled effect | 0.68 [0.49 – 0.87 |] |
| Baccarani, 2013 | PGWBI-DEP | 0.45 [-0.45 - 1.35 |] 5.6% |
| Kemeny, 2012 | | 0.43 [-0.43 - 1.33 | |
| | BDI | | - |
| Malarkey, 2013 | CES-D | | • |
| Manotas, 2014 | BSI-D | 0.62 [0.18 – 1.06 | - |
| Moody, 2013 | BDI | 0.15 [-0.47 - 0.77 | - |
| Pipe, 2009 | SCL-90-R-D | ⊢───────────────────────────────────── | - |
| Roeser, 2013 | BDI | ┝───━───┥ 0.51 [−0.04 − 1.06 | - |
| Wolever, 2012 | CES-D | 0.24 [-0.10 - 0.58 |] 17.8% |
| Anviotu | Pooled effect | 0.38 [0.14 – 0.62 |] |
| Anxiety Baccarani, 2013 | | 0.56 [-0.34 - 1.46 | 1 44.40 |
| | PGWBI_ANX | | |
| Manotas, 2014 | BSI-ANX | 0.75 [0.27 – 1.23 | |
| Pipe, 2009 | SCL-90-R-A | ↓ 0.71 [−0.02 − 1.44 | |
| Roeser, 2013 | STAI | |] 33.3% |
| | Pooled effect | 0.62 [0.32 – 0.92 | 1 |
| Burnout | | · · · · · · · · · · · · · · · · · · · | _ |
| Ancona, 2014 | MBI–EE | 0.42 [-0.20 - 1.04 | |
| Flook, 2013 | MBI–EE | 1.28 [0.26 – 2.30 | - |
| Mackenzie, 2006 | MBI–EE | · · · · · · · · · · · · · · · · · · · |] 24.79 |
| Moody, 2013 | MBI-EE | -0.35 [-0.94 - 0.24 |] 27.9% |
| Flook, 2013 | MBI-PA | <mark> ; −0.07 − 1.8</mark> 9 |] 21.2% |
| Mackenzie, 2006 | MBI-PA | h 0.69 [-0.04 - 1.42 |] 33.4% |
| Moody, 2013 | MBI-PA | 0.09 [-0.50 - 0.68 | 45.49 |
| Flook, 2013 | MBI-DP | 0.03 [-0.91 - 0.97 | |
| Mackenzie, 2006 | MBI-DP | 0.31 [-0.42 - 1.04 | |
| Moody, 2013 | MBI-DP | 0.12 [-0.47 - 0.71 | - |
| 1100003, 2010 | Pooled effect | 0.37 [0.07 – 0.67 | - |
| Wellbeing | | • | - |
| Baccarani, 2013 | PGWBI–TGWB | ⊢ <mark>⊢−−−−−−−−−−−−−−−−−−−−−−−−−−−−−−−−−−−</mark> | |
| McConachie, 2014 | WEMWS | ⊢ | - |
| Prasek, 2015 | WHO-5 | 0.42 [0.08 – 0.76 |] 16.2% |
| Bartlett, 2016 | AQoL–4D | 0.77 [0.25 – 1.29 |] 12.3% |
| Van Berkel, 2014 | SF-36 | ⊢− 0.07 [−0.35 − 0.21 |] 17.6% |
| Mackenzie, 2006 | SWLS | 0.60 [-0.13 - 1.33 | - |
| Shapiro, 2005 | SWLS | — 0.77 [–0.04 – 1.58 | - |
| Crain, 2016 | Work-Life | 0.77 [0.38 – 1.16 | |
| J. M.I., 2010 | Pooled effect | | - |
| Sleep | Fooled effect | 0.45 [0.17 – 0.72 | 1 |
| Bartlett, 2016 | JSQ | ⊢−−−− 0.45 [−0.07 − 0.97 |] 10.8% |
| Crain, 2016 | QoS | 0.13 [-0.26 - 0.52 |] 18.9% |
| Klatt, 2009 | PSQI | 0.27 [-0.35 - 0.89 | - |
| Malarkey, 2013 | PSQI | 0.28 [0.00 - 0.56 | - |
| Wolever, 2012 | PSQI | | - |
| , | | • 0.26 [0.09 – 0.43 | - |
| | Pooled effect | •.20 [0.09 - 0.43 | 1 |
| | | | |
| | | | |
| | | · · · · | |

Figure 4 Legend: K-10: Kessler 10-item Measure of Psychological Distress; SCL-90-R: Symptom Checklist-90-Revised (GSI = Global Severity Index, D = Depression, A = Anxiety); PDMS: Psychological Distress Manifestation Scale; GHQ-12: General Health Questionnaire (CHQ-12 = Chinese version); BSI-GSI: Brief Symptom Inventory: Global Severity Index; PGWBI: Psychological General Wellbeing Index (TGWB = Total General Wellbeing, DEP = Depression, ANX = Anxiety); STAI: State Trait Anxiety Index; MBE: Maslach Burnout Inventory (EE = Emotional Exhaustion, PA = Personal Accomplishment; DP = Depresonalisation); WEMWS: Wawrick-Edinburgh Mental Well-being Scale; WHO-5: World Health Organisation Wellbeing Index; AQoL-4D: Assessment of Health Related Quality of Life – 4 Dimensions; SF-36: Short Form Health Survey; SWLS: Satisfaction with Life Scale; JSQ: Jenkins Sleep Questionnaire; QoS: Quality of Sleep Questionnaire; PSQI: Pittsburgh Sleep Quality Index